Bernhard Nebel, Felix Lindner, Thorsten Engesser, Barbara Kuhnert, Laura Wächter WS 2017/18



Non-parametric Tests



- Wilcoxon signed-rank test
- Wilcoxon rank-sum test (aka Mann-Whitney test)
- Kruskal-Wallis test

Ranks



- Ranks are natural numbers starting with 1, which get assigned to scores sorted in increasing order.
- Ranks can be assigned to any data which is at least ordinal.
- Ranks are robust against outliers (because ranks are used instead of the actual data).

Example

■ Data: 0, 7, 3; Rank: 1, 3, 2

■ Data: -100, 99, 98; Rank: 1, 3, 2

■ Data: d, a, b; Rank: 3, 1, 2

Ranks: Ties



In case of ties, the average rank is assigned to the whole group of scores that constitutes the tie.

Example

■ Data: 1, 6, 4, 4, 2, 2, 2

■ Rank: 1, 7, 5.5, 5.5, 3, 3, 3

- Likert scales are a popular means of measurement.
- Likert scales in most cases have no interval-scale reading.

Five participants are asked to rate their belief in the possibility that humans will one day be the slaves of robots before and after they have watched a Sci-Fi movie. As a measurement instrument, a 3-Point Likert-Scale "never ever!" (1), "maybe" (2), "yes, sure!" (3) was used.

■ Before: 1, 2, 2, 3, 3; After: 2, 3, 3, 3, 1

■ Difference: -1, -1, -1, 0, +2

Differences without 0: -1, -1, -1, +2

Ranks: 2, 2, 2, 4

Five participants are asked to rate their belief in the possibility that humans will one day be the slaves of robots before and after they have watched a Sci-Fi movie. As a measurement instrument, a 3-Point Likert-Scale "never ever!" (1), "maybe" (2), "yes, sure!" (3) was used.

- Before: 1, 2, 2, 3, 3; After: 2, 3, 3, 3, 1
- Difference: -1, -1, -1, 0, +2; Without 0: -1, -1, -1, +2
- Ranks: 2, 2, 2, 4
- Let $V = \sum_{i=1}^{n} Z_{i}R_{i}$ be the sum of the positive ranks ($Z_{i} = 1$ if difference i is positive, and $Z_{i} = 0$ else).
- In the example V = 4. Well, so what?



- Imagine two paired samples and consider their rank differences.
- Consider $V = \sum_{i=1}^{n} Z_{i}R_{i}$. What could happen?
 - Case V = 0: All the rank differences are negative.
 - Case $V = \sum_{i=1}^{n} R_i = \frac{n(n+1)}{2}$: All rank differences are positive.
 - Blse: V ranges between 0 and $\frac{n(n+1)}{2}$.
- If the groups do not differ (H_0) , then 50% of the differences should be below 0 and 50% above. This is like saying that the median of the difference is 0. And in that case, V should be close to $\frac{n(n+1)}{2} = \frac{n(n+1)}{4}$.
- Hence, we will test H_0 : Mdn = 0 against its alternatives, and we will do that by using V.

■ The nice thing about V is that (for n > 25) its distribution is well approximated by a normal distribution $\mathcal{N}(\mu_V, \sigma_V)$ with

The nice thing about V is that (for n > 25) its distribution under H_0 is well approximated by a normal distribution $\mathcal{N}(\mu_V, \sigma_V)$ with

$$\mu_V = \frac{n(n+1)}{4}$$

$$\sigma_V = \sqrt{\frac{n(n+1)(2n+1)}{24}}$$

Proof (Mean): We already came to this conclusion earlier on Slide 8.



- The nice thing about V is that (for n > 25) its distribution is well approximated by a normal distribution $\mathcal{N}(\mu_V, \sigma_V)$ with
 - $\mu_V = \frac{n(n+1)}{4}$ $\sigma_V = \sqrt{\frac{n(n+1)(2n+1)}{24}}$
- Proof (Variance)
 - First, we define $V' = \sum_{i=1}^{n} V'_{i}$ with $V_i' = \begin{cases} 0 & \text{with probability } 0.5\\ i & \text{with probability } 0.5 \end{cases}$

 - \blacksquare (V' has the same distribution as V, because, for every rank, it either belongs to the sum of V or not with probability 0.5.)
 - $Var(V) = Var(V') = \sum_{i=1}^{n} Var(V'_i)$ (independence of V'_i).
 - $Var(V_i') = E(V_i'^2) E(V_i')^2 = (0^2 \frac{1}{2} + i^2 \frac{1}{2}) (\frac{1}{2}i)^2 = \frac{i^2}{4}$
 - $Var(V) = \sum_{i=1}^{n} Var(V_i) = \sum_{i=1}^{n} \frac{i^2}{4} = \frac{n(n+1)(2n+1)}{24}$.

Five participants are asked to rate their belief in the possibility that humans will one day be the slaves of robots before and after they have watched a Sci-Fi movie. As a measurement instrument, a 3-Point Likert-Scale "never ever!" (1), "maybe" (2), "yes, sure!" (3) was used.

- Before: 1, 2, 2, 3, 3; After: 2, 3, 3, 3, 1
- Difference: -1, -1, -1, 0, +2; Without 0: -1, -1, -1, +2
- Ranks: 2, 2, 2, 4

$$V = 4$$
, $\mu_V = 4(4+1)/4 = 5$, $\sigma_V = \sqrt{4(4+1)(2\times 4+1)/24}$

$$z = \frac{V - \mu_V}{\sigma_V} = (4 - 5)/2.74 = -0.365$$

$$p = P(z \le -0.365) + 1 - P(z \le 0.365) = 0.715$$

Example: t-Test

Five participants are asked to rate their belief in the possibility that humans will one day be the slaves of robots before and after they have watched a Sci-Fi movie. As a measurement instrument, a 3-Point Likert-Scale "never ever!" (1), "maybe" (2), "yes, sure!" (3) was used.

- Before: 1, 2, 2, 3, 3; After: 2, 3, 3, 3, 1
- Difference: -1, -1, -1, 0, +2
- $D = 0.20, s_D = 1.30, n = 5$
- $t = \sqrt{5} \times 0.20/1.30 = 0.344$, df = 4
- p = 0.748

Five participants are asked to rate their belief in the possibility that humans will one day be the slaves of robots after they have watched the Sci-Fi movie M1, and five participants rate their belief after watching Sci-Fi movie M2. As a measurement instrument, a 3-Point Likert-Scale "never ever!" (1), "maybe" (2), "yes, sure!" (3) was used.

- M1: 1, 1, 2, 2, 2; M2: 2, 3, 3, 3, 2
- \blacksquare H_0 : The two groups are equal.
- Reject H₀ or not?

- First, all scores are ranked together.
- First group's rank sum: $R_1 = \sum_{i=1}^{n_1} r_{1,i}$
- Second group's rank sum: $R_2 = \sum_{i=1}^{n_2} r_{2,i}$
- First group's W: $W_1 = R_1 \sum_{i=1}^{n_1} i = R_1 \frac{n_1(n_1+1)}{2}$
- Second group's W: $W_2 = R_2 \sum_{i=1}^{n_2} i = R_2 \frac{n_2(n_2+1)}{2}$
- $W_1 + W_2 = R_1 \frac{n_1(n_1+1)}{2} + R_2 \frac{n_2(n_2+1)}{2} = n_1 n_2$
- Note: The Wilcoxon Rank-Sum Test is also known as Mann-Whitney U-Test, and W is also called U. There are various ways of defining W (resp. U), which are all equal! R uses the statistics W the way shown above.



- For larger samples ($n_1 > 10, n_2 > 10$), $W \sim \mathcal{N}(\mu_W, \sigma_W)$:
 - $\mu_W = \frac{n_1 n_2}{2}$ $\sigma_W = \sqrt{\frac{n_1 n_2 (n_1 + n_2 + 1)}{12}}$
 - Also see simulation in lecture11. Rmd in the git repository slides.
- Again, we can calculate z-values to decide whether or not W is extreme, i.e., whether or not to reject H_0 .

Wilcoxon Rank-Sum Test: Example Continued



Example

- M1: 1, 1, 2, 2, 2
- M2: 2, 3, 3, 3, 2
- All Scores: 1, 1, 2, 2, 2, 2, 3, 3, 3, 2
- Ranks: 1.5, 1.5, 5, 5, 5, 5, 9, 9, 9, 5

$$\blacksquare$$
 $R_1 = 18, W = 18 - 15 = 3$

$$Z = \frac{\frac{3 - (5 \times 5)}{2}}{\sqrt{\frac{5 \times 5(5 + 5 + 1)}{12}}} = -2.298$$

$$p = P(z \le -2.298) + 1 - P(z \le 2.298) = 0.022$$



■ M1: 1, 1, 2, 2, 2

■ M2: 2, 3, 3, 3, 2

$$\overline{X}_1 = 1.6, \overline{X}_2 = 2.6, s_1^2 = 0.3, s_2^2 = 0.3, n = 5, df = 8$$

$$p = P(t \le -2.887) + 1 - P(t \le 2.887) = 0.020$$

■ For a simulation comparing Wilcoxon and t-Test see lecture11.Rmd in the git repository.

- Also for rank-based methods, there is an analog to ANOVA that can cope with more than two groups: Kruskal-Wallis Test. As for ANOVA, H₀ reads "There is no difference between the groups".
- First, the scores of all groups are ranked together (like for Wilcoxon Rank-Sum Test).
- The test statistics is called H:

■
$$H = (N-1)\frac{\sum_{i=1}^{p}n_{i}(\bar{r}_{i}-\bar{r})^{2}}{\sum_{i=1}^{p}\sum_{j=1}^{n_{i}}(r_{ij}-\bar{r})^{2}}$$
, with $N = \sum_{i=1}^{p}n_{i}$, $\bar{r}_{i} = \frac{\sum_{j=1}^{n_{i}}r_{ij}}{n_{i}}$, $\bar{r}_{i} = \frac{N+1}{2}$

- H can be simplified to $H = \frac{12}{N(N+1)} \sum_{i=1}^{p} n_i \bar{r}_i^2 3(N+1)$
- $H \sim \chi_{p-1}^2$, with p being the number of groups.

- M1: 1, 1, 2, 2, 2; Ranks: 2.5, 2.5, 8.5, 8.5, 8.5
- M2: 2, 3, 3, 3, 2; Ranks: 8.5, 14, 14, 14, 8.5
- M3: 1, 2, 2, 1, 2; Ranks: 2.5, 8.5, 8.5, 2.5, 8.5
- $\overline{r}_1 = 6.1, \overline{r}_2 = 11.8, \overline{r}_3 = 6.1, N = 15, \overline{r} = (15+1)/2 = 8$
- $H = \frac{12}{15 \times 16} \times 5(37.21 + 139.24 + 37.21) 3 \times 16 = 5.41$
- $p = 1 P(\chi^2 \le 5.41) = 0.067$
- R will report different values, see next slide to learn why.

Ties call for Corrections



If there are long ties (i.e., a lot of scores getting the same rank), the variance of the statistics become smaller and thus some corrections have to be applied.

■ The V-statistics's standard deviation becomes:

■ The W-statistics's standard deviation becomes:

$$\sigma_W = \sqrt{\frac{n_1 n_2}{12} \left((n_1 + n_2 + 1) - \sum_{i}^{k} \frac{t_i^3 - t_i}{(n_1 + n_2)(n_1 + n_2 - 1)} \right)}$$
 (cf., slide 16)

- And the H-statistics can be corrected by dividing H by the term $corr = 1 \frac{\sum_{i}^{k} (t_{i}^{3} t_{i})}{N^{3} N}$
 - In the example: $corr = 1 \frac{(4^3-4)+(8^3-8)+(3^3-3)}{(15^3-15)}$
 - The corrected H value then is $H_{corr} = 6.56$
- Because all this is rather tedious, you are allowed to skip these corrections in your assignments (also in the exam).



- Categorical Scale
 - χ^2 -statistics (χ^2 -distributed)
- Interval Scale
 - Variance known: z-statistics (normally distributed)
 - Variance unknown (but equal): t-statistics (Student's t distribution), F-statistics (F-distributed)
- Ordinal Scale
 - W-, V-statistics (both normally distributed), H-statistics (χ^2 -distributed)



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- We started out defining four types of hypotheses
 - Directional difference hypotheses
 - 2 Undirectional difference hypotheses
 - 3 Directional relationship hypotheses
 - 4 Undirectional relationship hypotheses
- We can so far only deal with (1) and (2). This is going to be fixed during the next statistics block starting from January 16th. The timeline:
 - Next: We will have a mock exam on January 9th.
 - NextNext: We will have another Reading Session on January 11.
 - NextNextNext: Correlation and Regression for testing relationship hypotheses from January 16th.
 - New Exam Date: Feburary 21th!

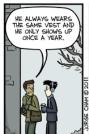
Merry Christmas!



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Sketches

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